# Assurances for machine learning trajectory predictors: guaranteed probabilistic bounds with conformal prediction

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This is mainly going to be an introductory talk in conformal prediction.
I will try to show you that it's a very simple yet powerful method.
I will introduce it in the context of my work.
I will also give you a glimpse of some results.

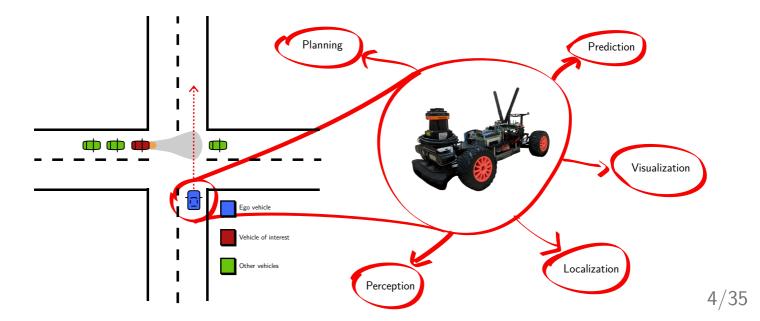
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## What do I do?



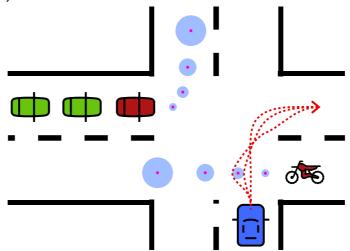
#### What do I do?

I develop planning algorithms taking **probabilistic motion predictions** of other traffic participants. These algorithms should be able to guarantee **safety**, while being **real-time feasible**.

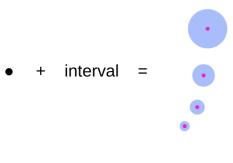


#### For this talk

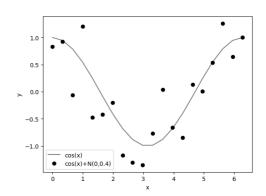
Given trajectory predictions in magenta, we want to compute valid prediction regions in blue, for a given desired coverage probability  $1-\alpha$  (probability true trajectory is inside the blue regions).



Given a dataset of 1-D points  $\mathcal{D} = \{(x_1, y_1), \dots, (x_n, y_n)\}$ , we want to compute valid prediction regions for a new point  $\hat{y}_{n+1} = f(x_{n+1})$ , for a given desired coverage probability  $1 - \alpha = 90\%$ .



$$y = \cos(x) + \mathcal{N}(0, 0.4)$$

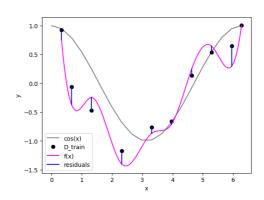


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We can train f using a subset of  $\mathcal{D}$ ,  $\mathcal{D}_{train} \subset \mathcal{D}$ .

How can we choose a band q, such that:

$$\mathbb{P}(y_{n+1} \in [f(x_{n+1}) - q, f(x_{n+1}) + q]) \ge 90\%$$



Given a dataset of 1-D points  $\mathcal{D} = \{(x_1, y_1), \dots, (x_n, y_n)\}$ , we want to compute valid prediction regions for a new point  $\hat{y}_{n+1} = f(x_{n+1})$ , for a given desired coverage probability  $1 - \alpha = 90\%$ .

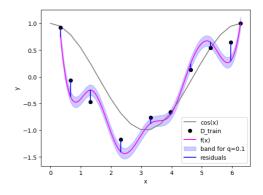
Given the residuals defined as:

$$r_i = |y_i - f(x_i)|$$

The problem is equivalent to finding a band q, such that:

$$\mathbb{P}(r_{n+1} \leq q) \geq 90\%$$

One possible solution is to use the 90% empirical quantile of the residuals.



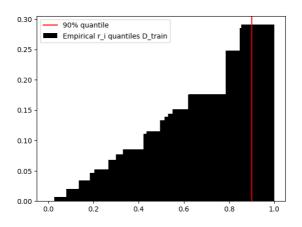
#### Brief recap on quantiles

Given a distribution F the level  $\beta$  quantile is defined as follows, for  $Z \sim F$ :

$$Quantile(\beta, F) = \inf\{z : \mathbb{P}(Z \leq z) \geq \beta\}$$

For an empirical distribution X (such as the residuals  $r_i$  on the dataset  $D_{train}$ ) it can be defined as:

$$Quantile(\beta, X) = Quantile(\beta, \frac{1}{n} \sum_{i=1}^{n} \delta_{x_i})$$

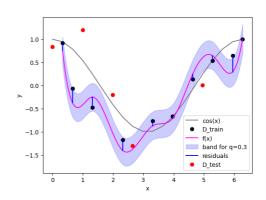


Given a dataset of 1-D points  $\mathcal{D} = \{(x_1, y_1), \dots, (x_n, y_n)\}$ , we want to compute valid prediction regions for a new point  $\hat{y}_{n+1} = f(x_{n+1})$ , for a given desired coverage probability  $1 - \alpha = 90\%$ .

Given the residuals defined as:

$$r_i = |y_i - f(x_i)|$$

Compute the q=90% empirical quantile of the residuals, for the points in  $\mathcal{D}_{train}$ , and take  $[f(x_{n+1})-q,f(x_{n+1})+q]$ .



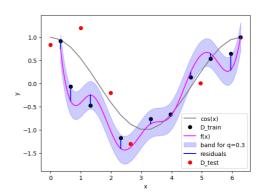
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Only 40% coverage on a test dataset  $\mathcal{D}_{test}$  disjoint with  $\mathcal{D}_{train}$  (only a sample of the test points is shown in the image)!



# A little bit of history

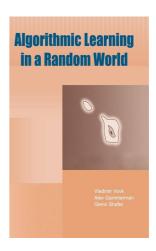


Vladimir Vovk

## A little bit of history



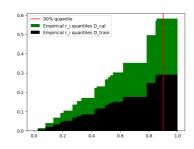
Vladimir Vovk

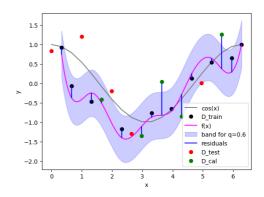


Algorithmic learning in a Random World

Given a dataset of 1-D points  $\mathcal{D} = \{(x_1, y_1), \dots, (x_n, y_n)\}$ , we want to compute valid prediction regions for a new point  $\hat{y}_{n+1} = f(x_{n+1})$ , for a given desired coverage probability  $1 - \alpha = 90\%$ .

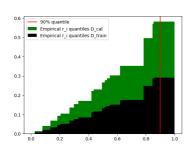
Compute the q=90% empirical quantile of the residuals, for the points in  $\mathcal{D}_{cal}$ , not used for the training of f! Then take  $[f(x_{n+1})-q,f(x_{n+1})+q]$ .

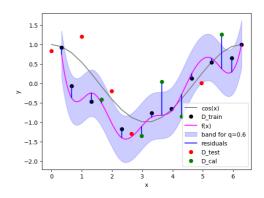




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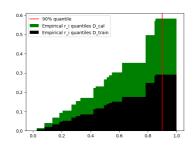


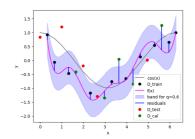


It works!

#### Split conformal prediction

- Compute the residuals  $r_i = |f(x_i) y_i|$  using the pairs  $(x_i, y_i)$  in the calibration dataset  $\mathcal{D}_{cal}$ .
- Sort the residuals in ascending order:  $r_{(1)} \leq r_{(2)} \leq \cdots \leq r_{(n)}$ .
- Given a max error probability of  $\alpha$ , select the  $q_{1-\alpha} = \lceil (n+1)(1-\alpha) \rceil$ -th residual in ascending order.
- The prediction set is the set of labels y such that  $|f(x) y| \le q_{1-\alpha}$ . More explicitly  $[f(x) q_{1-\alpha}, f(x) + q_{1-\alpha}]$ .





#### Split conformal prediction

- Take any measurable score function s(x, y) (residual was r = |f(x) y|, with f the predictor trained in  $\mathcal{D}_{train}$ ).
- Compute the  $1-\alpha$  quantile of the scores on the calibration dataset  $(Quantile(1-\alpha, \mathcal{D}_{cal}))$ .

$$\hat{C}(x) = \{ y \text{ s.t } s(x, y) \leq Quantile(1 - \alpha, \mathcal{D}_{cal}) \}$$

Theorem [Vovk, Gammerman, Shafer 2005]

$$\mathbb{P}\left\{Y\in\hat{C}(X)\right\}\geq 1-\alpha$$

Holds as long as the new data (X,Y) is exchangeable with the calibration dataset  $\mathcal{D}_{cal}$ .

## Split conformal prediction - Proof

Given a sequence of random variables:

$$R_1, R_2, \ldots, R_n, R_{n+1}, \ldots$$

Suppose that any permutation is equally likely. That is the sequence is exchangeable, more formally:

$$\mathbb{P}\left\{R_{1} \leq r_{1}, R_{n+1} \leq r_{n+1}, \dots\right\} = \mathbb{P}\left\{R_{\pi(1)} \leq r_{1}, R_{\pi(n+1)} \leq r_{n+1}, \dots\right\}$$

For all permutations,  $\pi$  and all  $r_i$ .

#### Split conformal prediction - Proof

This means that  $R_{n+1}$  is equally likely to be among the k smallest values among  $R_1, \ldots, R_{n+1}$ . Suppose that  $R_i$  are different almost surely this translates to:

$$\mathbb{P}\left\{R_{n+1} \text{ is among the } k \text{ smallest in } R_1,\ldots,R_{n+1}\right\} = \frac{k}{n+1}$$

Which is equivalent to:

$$\mathbb{P}\left\{R_{n+1} \text{ is among the } k \text{ smallest in } R_1,\ldots,R_n\right\} = \frac{k}{n+1}$$

Taking  $k = \lceil (n+1)(1-\alpha) \rceil$  we have:

$$\mathbb{P}\left\{R_{n+1} \text{ is among the } k \text{ smallest in } R_1,\ldots,R_n\right\} = \frac{\lceil (n+1)(1-\alpha) \rceil}{n+1}$$

## Split conformal prediction - Proof

We have:

$$\mathbb{P}\left\{R_{n+1} \text{ is among the } k \text{ smallest in } R_1,\ldots,R_{n+1}\right\} \in [1-\alpha,1-\alpha+1/(n+1))$$

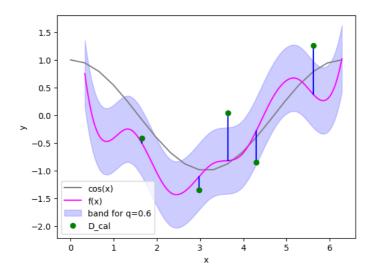
We can translate what is inside  $\mathbb{P}$  to:

$$q = Quantile\left(\frac{\lceil (n+1)(1-lpha)
ceil}{n}, \frac{1}{n}\sum_{i=1}^n \delta_{R_i}
ight)$$

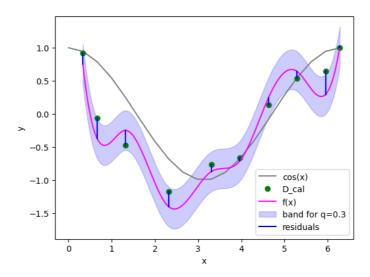
Finally, we can write:

$$1-\alpha \leq \mathbb{P}\left\{R_{n+1} \leq q\right\} < 1-\alpha + \frac{1}{n+1}$$

Does that mean that we can use any calibration dataset size? What happens if we use a very small calibration dataset?



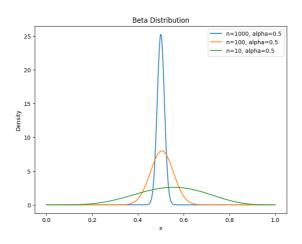
Does that mean that we can use any calibration dataset size? What happens if we use a very small calibration dataset?



Conditional on the data of the calibration dataset, the coverage is distributed as:

$$\mathbb{P}\left\{Y\in \hat{\mathcal{C}}(X) \mid (X_i,Y_i)\in \mathcal{D}_{cal}
ight\}\sim extit{Beta}(k,n-k+1)$$

With  $k = \lceil (n+1)(1-\alpha) \rceil$ .

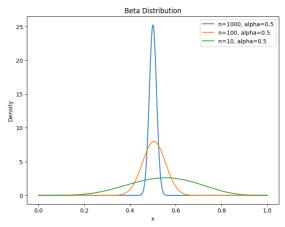


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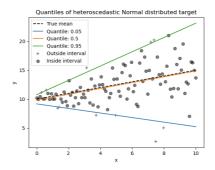
With  $k = \lceil (n+1)(1-\alpha) \rceil$ .

n (size of $D_{cal}$ )	coverage correct $+/-5\%$	
10	0.24	
100	0.68	
1000	0.99	



## Conformalized quantile regression

In the same spirit, we can try to find adaptive bounds. One possibility is to train our predictor f to output quantiles as it is done with the quantile regression:



In this case, the outputs of our predictor are given by  $f(x)=\{f_{\frac{\alpha}{2}}(x),f_{1-\frac{\alpha}{2}}(x)\}$ 

## Conformalized quantile regression

This can be easily achieved for any learning based predictor by just using the pinball loss:

$$\mathcal{L}_{\alpha}(y, f(x)) = \begin{cases} \alpha(y - f(x)) & \text{if } y > f(x) \\ (1 - \alpha)(f(x) - y) & \text{otherwise} \end{cases}$$

Using the following conformity score:

$$s(x,y) = \max\left\{y - f_{\frac{\alpha}{2}}(x), f_{1-\frac{\alpha}{2}}(x) - y\right\}$$

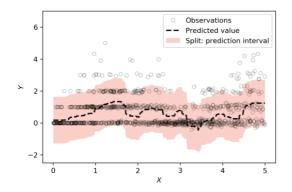
We can, now compute the quantile q and build our conformal predictor.

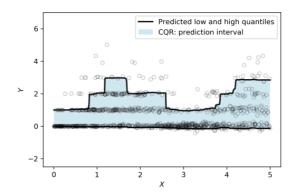
## Conformalized quantile regression

Our conformal bands will be given by:

$$\hat{C}(x) = \left[ f_{\frac{\alpha}{2}}(x) - q, f_{1-\frac{\alpha}{2}}(x) + q \right]$$

This method is introduced by *Romano et al. (2019)*. Here are some of their showcase results:





Conformal prediction for time series was introduced by *Stankeviciute et al. (2021)*. Some other work has followed the same idea as well.

Given a discrete time series of length n, for simplicity, of real values:

$$(y_1,\ldots,y_n)$$

We want to predict the values from  $y_{m+1}$  to  $y_n$ , given the values from  $y_1$  to  $y_m$ . This is done via a neural network with m inputs and n-m outputs:

$$f(y_1, ..., y_m) = (\hat{y}_{m+1}, ..., \hat{y}_n)$$

The idea is to build one conformal predictor for each output, of the neural network, independently. Our conformal predictor will look like:

$$\hat{C}(y_1,\ldots,y_m)=\left(\hat{C}_{m+1}(y_1,\ldots,y_m),\ldots,\hat{C}_n(y_1,\ldots,y_m)\right)$$

For a coverage of  $1-\beta$  for each individual predictor, each will have an error probability of  $\beta$ . Therefore the probability of at least one error obeys the following given Boole's inequality:

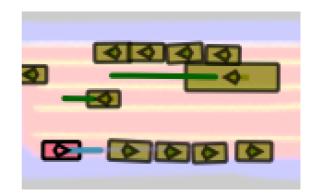
$$\mathbb{P}\left\{\bigcup_{i=m+1}^n \hat{y}_i \notin \hat{C}_i(y_1,\ldots,y_m)\right\} \leq \sum_{i=m+1}^n \mathbb{P}\left\{\hat{y}_i \notin \hat{C}_i(y_1,\ldots,y_m)\right\}$$

The probability of at least one error is at least  $\beta(n-m)$ . So if we want all predictions to be valid (tube around our prediction) we want to choose  $\beta = \frac{\alpha}{(n-m)}$ .

Applying this idea to Trajectron++ with a time step of 0.5s, we have:

Table 1: Prediction set sizes for a  $1-\alpha=90\%$ 

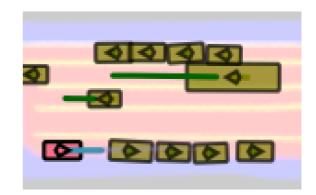
x(n	1)	y(m)	t(s)
0.88	86	0.1881	0.5
1.59	65	0.3520	1.0
2.22	46	0.5275	1.5
3.08	81	0.7505	2.0
4.22	29	1.0411	2.5



If full coverage is not required Trajectron++ with a time step of 0.5s, we have:

Table 2: Prediction set sizes for a  $1 - \alpha = 90\%$ 

x(m)	y(m)	t(s)
0.5345	0.0711	0.5
0.7267	0.1453	1.0
1.0552	0.2060	1.5
1.5524	0.2999	2.0
2.1262	0.3672	2.5



#### Some stuff I did not cover

- Conformal prediction for classification : Adaptive prediction sets.
- Use part of calibration data for training: Full conformal prediction, Cross-Conformal Prediction, CV+, and Jackknife+.
- Online updates: Rolling RC and adaptive conformal prediction.
- Conformal prediction when we face distribution shifts: *Conformal prediction under the covariate shift*.

#### Conclusions

#### Takeaways:

- Easy to implement and efficient.
- Provides valid guarantees with finite samples.
- Active area of research, lots of new papers per year.
- Used in world scenarios.

#### Be attentive to:

- Distribution shifts or anything that breaks the exchangeability assumption.
- Conditional validity could be a problem.

#### References I

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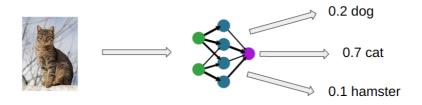
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- Ryantibs/Statlearn-S23: Course Materials for Advanced Topics in Statistical Learning, Spring 2023. URL: https://github.com/ryantibs/statlearn-s23/tree/main (visited on 11/13/2023).

## What about classification problems?

Given a set of labels  $\mathcal{Y} = \{cat, dog, hamster\}$ , neural networks are able to output estimates of their likelihoods  $f(x) = \{\hat{p}_{cat}, \hat{p}_{dog}, \hat{p}_{hamster}\}$ :



How to provide safe prediction sets in such scenarios?

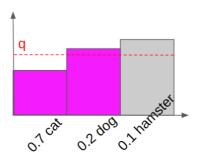
#### Adaptive prediction sets

APS - adaptive prediction sets (Romano et al. (2020) Angelopoulos et al. (2021)) solves this issue.

We define  $\pi$  to be a permutation which sorts the outputs of the predictor  $f(x) = \{\hat{p}_1, \dots, \hat{p}_K\}$  in decreasing order. Our conformal predictor will be:

$$\hat{C}(x) = \left\{\hat{p}_{\pi(1)}, \dots, \hat{p}_{\pi(k)}\right\}$$

Where k is chosen such that we have a cumulative sum until we reach the quantile q over the conformity scores corresponding to  $1-\alpha$  coverage:



#### Adaptive prediction sets

The quantile q is chosen as previously:

$$q = Quantile\left(\frac{\lceil (n+1)(1-lpha) \rceil}{n}, \frac{1}{n} \sum_{i=1}^{n} \delta_{s_i}\right)$$

The conformity scores  $s_i$  are defined as:

$$s(x,y) = \sum_{j=1}^k \hat{p}_{\pi(j)}$$
 where  $y = \pi(k)$ 

## How to make the coverage correct per class

We wish that, in the classification setting, we could have the coverage guarantees per class, more formally:

$$\mathbb{P}\left\{Y\in\hat{C}(X)\mid Y=y\right\}\geq 1-\alpha$$

If  $\mathcal{Y} = \{Sick, Healthy\}$ , we would like to have prediction sets valid independent of the true label.

	Sick	Healthy
Test Positive	60%	40%
Test Negative	10%	90%

#### How to make the coverage correct per class

If we define the quantiles per class as:

$$q^k = extit{Quantile}\left(rac{\lceil (n^k+1)(1-lpha)
ceil}{n^k}, rac{1}{n}\sum_{i=1}^{n^k}\delta_{s_i^k}
ight)$$

Where the superscript k denotes restricts the samples in  $\mathcal{D}_{cal}$  to the class k. We can define a class conditional valid conformal predictor as:

$$\hat{C}(x) = \{ y \text{ s.t } s(x, y) \leq q^y \}$$

#### What else could we want?

Instance conditional validity:

$$\mathbb{P}\left\{Y\in\hat{C}(X)\mid X=x\right\}\geq 1-\alpha$$

Unfortunately, that's impossible :(

**Lei and Wasserman (2014)** ...any prediction band which claims to cover at almost every point, for every joint distribution, must be infinite in size ...

#### But not everything is lost

Group conditional validity:

$$\mathbb{P}\left\{Y\in\hat{C}(X)\mid X\in\mathcal{G}_i\right\}\geq 1-\alpha$$

Given a partition of the input space  $\mathcal{G}_1, \dots, \mathcal{G}_k$ . We can define a group conditional valid conformal predictor as:

$$\hat{C}(x) = \{ y \text{ s.t } s(x,y) \le q^g \}$$

